

Conditional Expectation

1. Remember that conditional expectation $E[Y|\mathbf{X}]$ is a function of \mathbf{X} . When confused, remember it is in essence an expression $\mu(\mathbf{x}) = E[Y|\mathbf{X} = \mathbf{x}]$ for all possible \mathbf{x} , and then we replace $\mathbf{X} = \mathbf{x}$ with \mathbf{X} (this is just for simplicity of notation) to get $\mu(\mathbf{X}) = E[Y|\mathbf{X}]$.
2. $E[Y|\mathbf{X}]$ is the projection of Y on \mathbf{X} , i.e., it minimizes mean squared error:

$$E[Y|\mathbf{X}] = \arg \min_{g(\mathbf{X})} E[(Y - g(\mathbf{X}))^2]$$

3. CEF error is orthogonal to every function of \mathbf{X} , i.e., for all functions h , we have

$$E[(Y - E[Y|\mathbf{X}])h(\mathbf{X})] = 0$$

A quick proof using the Law of iterated expectation and factoring out $h(X)$:

$$\begin{aligned} E[(Y - E[Y|X])h(X)] &= E[h(X)Y] - E[E[Y|X]h(X)] \\ &= E[E[h(X)Y|X]] - E[E[Y|X]h(X)] \\ &= E[h(X)E[Y|X]] - E[E[Y|X]h(X)] = 0 \end{aligned}$$

4. Important properties of conditional expectation:

- (a) If $X \perp Y$, we have $E[Y|X] = E[Y]$, $E[XY] = E[X]E[Y]$.
- (b) $E[h(X)|X] = h(X)$. Notice that the result is still a function of X rather than an expected value. $E[h(X)Y|X] = h(X)E[Y|X]$ for all functions h .
- (c) Linearity: $E[cY|X] = cE[Y|X]$ for any constant c .
- (d) Law of iterated expectation (Adam's Law, law of total expectation): $E[Y] = E[E[Y|X]] = \int_x E[Y|X = x]f_X(x) dx = E[Y|B]P(B) + E[Y|B^c]P(B^c)$. With extra conditioning, we have $E[Y|Z] = E[E[Y|X, Z]|Z] = \int_x E[Y|X, Z]f_{X|Z}(x|Z) dx$.

5. Conditional variance $\text{Var}(Y|\mathbf{X})$ is also a function of \mathbf{X} , denote as $\sigma^2(\mathbf{X})$. We have

$$\text{Var}(Y|X) = E[(Y - E[Y|X])^2|X] = E[Y^2|X] - (E[Y|X])^2$$

which is repeatedly used in the proof of EVVE.

6. EVVE's law: $\text{Var}(Y|X) = E[\text{Var}(Y|X)] + \text{Var}(E[Y|X])$. We can interpret it as the sum of within-group variation and between-group variation.

Example

You are approached by a mysterious stranger, who allows you to bid on a mystery box containing a mystery prize! The value V of the prize is Uniform on $[0, 1]$ (measured in millions of dollars). You can choose to bid any amount b (in millions of dollars). If $b \geq 2V/3$ then your bid is accepted (and your payoff is $V - b$, value minus bid). Otherwise, the bid is rejected and nothing happens. What is your optimal bid?

Solution: Let W be the payoff. If $b < 2/3$, we have by law of total expectation,

$$\begin{aligned} E[W] &= E[W|b \geq 2V/3]P(b \geq 2V/3) + E[W|b < 2V/3]P(b < 2V/3) \\ &= E[V - b|b \geq 2V/3]P(b \geq 2V/3) + 0 \\ &= (E[V|V \leq 3b/2] - b)P(V \leq 3b/2) \\ &= \left(\int_0^{3b/2} \frac{2v}{3b} dv - b \right) \left(\int_0^{3b/2} 1 dv \right) \\ &= (3b/4 - b)(3b/2) = -3b^2/8 \leq 0 \end{aligned}$$

If $b \geq 2/3$, we always have $b \geq 2V/3$, and

$$E[W] = E[V - b] = \frac{1}{2} - b \leq -\frac{1}{6}$$

So the optimal bid is just 0.

Question 1

Suppose we want to model the relationship between legislation and politician quality. There are two types of politician quality: high and low. When a high quality politician proposes a bill, it has a probability p_1 to pass; conversely, when a low quality politician proposes a bill, it has a probability p_2 to pass, where $p_1 > p_2$. Unfortunately, we cannot directly observe politicians' quality, but instead rely on our prior that a politician is a high type with probability h and low type with probability $(1 - h)$, where $h \in (0, 1)$. Let X be the number of passed bills after a randomly picked politician has made n proposals.

- (a) Find the marginal distribution of X .
- (b) Find the mean and variance of X .

Solution to Question 1

(a) Marginal distribution of X . Let H denote the event that the politician is of high quality, and L the event of low quality.

$$X | H \sim \text{Binomial}(n, p_1), \quad X | L \sim \text{Binomial}(n, p_2).$$

By the law of total probability and using the prior that the politician is high quality with probability h ,

$$P(X = k) = P(X = k | H) P(H) + P(X = k | L) P(L).$$

Hence, the marginal distribution is a mixture of two Binomial distributions:

$$P(X = k) = h \binom{n}{k} p_1^k (1 - p_1)^{n-k} + (1 - h) \binom{n}{k} p_2^k (1 - p_2)^{n-k}, \quad k = 0, 1, \dots, n.$$

(b) Mean and variance of X .

$$\mathbb{E}[X] = \mathbb{E}[\mathbb{E}[X | \text{type}]] = h \mathbb{E}[X | H] + (1 - h) \mathbb{E}[X | L].$$

Since $X | H \sim \text{Binomial}(n, p_1)$ and $X | L \sim \text{Binomial}(n, p_2)$,

$$\mathbb{E}[X | H] = n p_1, \quad \mathbb{E}[X | L] = n p_2.$$

Therefore,

$$\mathbb{E}[X] = h n p_1 + (1 - h) n p_2 = n [h p_1 + (1 - h) p_2].$$

For the variance, use the law of total variance:

$$\text{Var}(X) = \mathbb{E}[\text{Var}(X | \text{type})] + \text{Var}(\mathbb{E}[X | \text{type}]).$$

We know

$$\text{Var}(X | H) = n p_1 (1 - p_1), \quad \text{Var}(X | L) = n p_2 (1 - p_2).$$

Hence,

$$\mathbb{E}[\text{Var}(X | \text{type})] = h (n p_1 (1 - p_1)) + (1 - h) (n p_2 (1 - p_2)).$$

Also,

$$\mathbb{E}[X | H] = n p_1, \quad \mathbb{E}[X | L] = n p_2.$$

So

$$\text{Var}(\mathbb{E}[X | \text{type}]) = h (n p_1)^2 + (1 - h) (n p_2)^2 - \left(n [h p_1 + (1 - h) p_2] \right)^2.$$

Putting it all together gives the total variance:

$$\text{Var}(X) = h n p_1 (1 - p_1) + (1 - h) n p_2 (1 - p_2) + h (n p_1)^2 + (1 - h) (n p_2)^2 - \left(n [h p_1 + (1 - h) p_2] \right)^2.$$

You can leave it in this form or simplify it as needed.

!!!Notice: If you pick a new random politician each time for a new proposal, then we can do $X \sim \text{Bin}(n, p_1 h + p_2 (1 - h))$ and directly derive the mean and variance.

Question 2

We know from the definition of the variance that

$$\mathbb{E}[(Y - \mathbb{E}[Y])^2] = \mathbb{E}[Y^2] - (\mathbb{E}[Y])^2.$$

Prove that this equality still holds when we condition on X , i.e.,

$$\mathbb{E}[(Y - \mathbb{E}[Y | X])^2 | X] = \mathbb{E}[Y^2 | X] - (\mathbb{E}[Y | X])^2.$$

Solution to Question 2

Recall that conditioning on X , we treat X as fixed while taking expectation with respect to Y . By definition:

$$\mathbb{E}[(Y - \mathbb{E}[Y | X])^2 | X] = \mathbb{E}[Y^2 - 2Y\mathbb{E}[Y | X] + (\mathbb{E}[Y | X])^2 | X].$$

Since $\mathbb{E}[Y | X]$ is treated as a constant when conditioning on X ,

$$= \mathbb{E}[Y^2 | X] - 2\mathbb{E}[Y | X]\mathbb{E}[Y | X] + (\mathbb{E}[Y | X])^2 = \mathbb{E}[Y^2 | X] - (\mathbb{E}[Y | X])^2.$$

Hence,

$$\mathbb{E}[(Y - \mathbb{E}[Y | X])^2 | X] = \mathbb{E}[Y^2 | X] - (\mathbb{E}[Y | X])^2,$$

as required.

Question 3

Let X_1, X_2, \dots, X_n be i.i.d. random variables with mean μ and variance σ^2 , and $n \geq 2$. A bootstrap sample of X_1, \dots, X_n is a sample of n random variables X_1^*, \dots, X_n^* formed from the X_j by sampling *with replacement* with equal probabilities. Let \bar{X}^* denote the sample mean of the bootstrap sample:

$$\bar{X}^* = \frac{1}{n}(X_1^* + \dots + X_n^*).$$

- (a) Find $\mathbb{E}[X_j^*]$ and $\text{Var}(X_j^*)$ for each j . (Hint: What is the distribution of X_j^* ?)
- (b) Find $\mathbb{E}[\bar{X}^* | X_1, \dots, X_n]$ and $\text{Var}[\bar{X}^* | X_1, \dots, X_n]$. (Hint: Conditional on X_1, \dots, X_n , the X_j^* are independent, each putting probability $1/n$ at each of the points X_1, \dots, X_n .)
- (c) Find $\mathbb{E}[\bar{X}^*]$ and $\text{Var}[\bar{X}^*]$. (Hint: Recall that the sample variance $\frac{1}{n-1} \sum_{j=1}^n (X_j - \bar{X})^2$ is an unbiased estimator of the population variance σ^2 .)

Solution to Question 3

Let us denote the original sample as X_1, \dots, X_n (i.i.d. with mean μ and variance σ^2). A bootstrap sample (X_1^*, \dots, X_n^*) is drawn *with replacement* from $\{X_1, \dots, X_n\}$.

(a) Distribution of each X_j^* . Conditioned on X_1, \dots, X_n , the random variable X_j^* is equally likely to be any of X_1, \dots, X_n , with probability $1/n$ each. Hence:

$$\mathbb{E}[X_j^* | X_1, \dots, X_n] = \frac{1}{n} \sum_{i=1}^n X_i.$$

Unconditionally, this random variable still has expectation μ , but more precisely,

$$\mathbb{E}[X_j^*] = \mathbb{E}[\mathbb{E}[X_j^* | X_1, \dots, X_n]] = \mathbb{E}\left[\frac{1}{n} \sum_{i=1}^n X_i\right] = \frac{1}{n} \sum_{i=1}^n \mathbb{E}[X_i] = \mu.$$

For the variance, again conditioning on X_1, \dots, X_n :

$$\text{Var}(X_j^* | X_1, \dots, X_n) = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - (\bar{X})^2.$$

By EVVE's law, we have

$$\begin{aligned} \text{Var}(X_j^*) &= \mathbb{E}[\text{Var}(X_j^* | X_1, \dots, X_n)] + \text{Var}(\mathbb{E}[X_j^* | X_1, \dots, X_n]) \\ &= \frac{1}{n} \sum_{i=1}^n \mathbb{E}[X_i^2] - \mathbb{E}[\bar{X}^2] + \text{Var}(\bar{X}) \\ &= \mathbb{E}[X_1^2] - (\mathbb{E}[X_1])^2 \\ &= \text{Var}(X_1) = \sigma^2 \end{aligned}$$

(b) **Mean and variance of \bar{X}^* conditional on X_1, \dots, X_n .** We have

$$\bar{X}^* = \frac{1}{n} \sum_{j=1}^n X_j^*.$$

Since, given X_1, \dots, X_n , each X_j^* is an i.i.d. draw from the empirical distribution that places mass $1/n$ on each X_i ,

$$\mathbb{E}[\bar{X}^* | X_1, \dots, X_n] = \frac{1}{n} \sum_{j=1}^n \mathbb{E}[X_j^* | X_1, \dots, X_n] = \frac{1}{n} \sum_{j=1}^n \bar{X} = \bar{X}.$$

Hence, conditional on the original sample, the mean of the bootstrap sample average is simply \bar{X} .

For the variance,

$$\text{Var}(\bar{X}^* | X_1, \dots, X_n) = \frac{1}{n^2} \sum_{j=1}^n \text{Var}(X_j^* | X_1, \dots, X_n),$$

because the X_j^* are independent given X_1, \dots, X_n . Each term $\text{Var}(X_j^* | X_1, \dots, X_n)$ is the same, so

$$= \frac{1}{n^2} n \text{Var}(X_1^* | X_1, \dots, X_n) = \frac{1}{n} \text{Var}(X_1^* | X_1, \dots, X_n).$$

Using the fact that, conditionally, X_1^* takes the values X_1, \dots, X_n each with probability $1/n$:

$$\text{Var}(X_1^* | X_1, \dots, X_n) = \frac{1}{n} \sum_{i=1}^n X_i^2 - (\bar{X})^2.$$

Thus

$$\text{Var}(\bar{X}^* | X_1, \dots, X_n) = \frac{1}{n} \left[\frac{1}{n} \sum_{i=1}^n X_i^2 - (\bar{X})^2 \right].$$

(c) **Mean and variance of \bar{X}^* unconditional.** First, for the mean:

$$\mathbb{E}[\bar{X}^*] = \mathbb{E}[\mathbb{E}[\bar{X}^* | X_1, \dots, X_n]] = \mathbb{E}[\bar{X}] = \mu.$$

Next, for the variance, use the law of total variance:

$$\text{Var}(\bar{X}^*) = \mathbb{E}[\text{Var}(\bar{X}^* | X_1, \dots, X_n)] + \text{Var}(\mathbb{E}[\bar{X}^* | X_1, \dots, X_n]).$$

From part (b),

$$\mathbb{E}[\bar{X}^* | X_1, \dots, X_n] = \bar{X}.$$

Hence

$$\text{Var}(\mathbb{E}[\bar{X}^* | X_1, \dots, X_n]) = \text{Var}(\bar{X}) = \frac{\sigma^2}{n}.$$

Also from part (b),

$$\text{Var}(\bar{X}^* | X_1, \dots, X_n) = \frac{1}{n} \left[\frac{1}{n} \sum_{i=1}^n X_i^2 - (\bar{X})^2 \right].$$

Taking expectation over (X_1, \dots, X_n) yields

$$\mathbb{E}[\text{Var}(\bar{X}^* | X_1, \dots, X_n)] = \frac{1}{n} \mathbb{E} \left[\frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2 \right].$$

Observe that

$$\mathbb{E} \left[\frac{1}{n} \sum_{i=1}^n X_i^2 \right] = \mathbb{E}[X^2] = \sigma^2 + \mu^2,$$

and

$$\mathbb{E}[\bar{X}^2] = \text{Var}(\bar{X}) + (\mathbb{E}[\bar{X}])^2 = \frac{\sigma^2}{n} + \mu^2.$$

Hence

$$\mathbb{E} \left[\frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2 \right] = (\sigma^2 + \mu^2) - \left(\frac{\sigma^2}{n} + \mu^2 \right) = \sigma^2 - \frac{\sigma^2}{n} = \sigma^2 \left(1 - \frac{1}{n} \right).$$

Thus

$$\mathbb{E}[\text{Var}(\bar{X}^* | X_1, \dots, X_n)] = \frac{1}{n} \sigma^2 \left(1 - \frac{1}{n} \right) = \frac{\sigma^2}{n} \left(1 - \frac{1}{n} \right) = \sigma^2 \frac{n-1}{n^2}.$$

Putting both pieces together for the total variance,

$$\text{Var}(\bar{X}^*) = \sigma^2 \frac{n-1}{n^2} + \frac{\sigma^2}{n} = \frac{\sigma^2(n-1)}{n^2} + \frac{\sigma^2}{n} = \frac{\sigma^2(n-1)}{n^2} + \frac{\sigma^2 n}{n^2} = \frac{\sigma^2(n-1+n)}{n^2} = \frac{\sigma^2(2n-1)}{n^2}.$$