

GOV 51 Section

Week 4: Introduction to Regression

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Prediction

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 - ▶ **Intercept** (β_0): average value of Y when X is 0
 - ▶ **Slope** (β_1): average change in Y when X increases by one unit.

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- ▶ **Regression line:** $\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 x$
 - ▶ Notice how we **NO LONGER** use the notation Y_i or $X_i \rightsquigarrow$ we are now thinking about averages.
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$$\mathbb{E}[Y | x] = \hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 x$$

Regression is not magic

- ▶ People love to regress some variable on another, but that doesn't inherently mean anything. If you regress some continuous variable (think income) on some binary variable (think treated vs. not treated) your β_0 and β_1 are just describing the average value within the treated group ($x = 0$) and the difference in the average values in the treated and non-treated group. Why?

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$$T_i \perp \{Y_i(1), Y_i(0)\}_{i=1}^n \mid \mathbf{X}_i \forall x_i$$

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- ▶ Finds the line that minimizes the magnitude of the prediction errors!
- ▶ Least squares line **always** goes through (\bar{X}, \bar{Y}) !

More on Regression & Model fit

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- ▶ Benchmarking our predictions using the **proportional reduction in error**:

$$\frac{\text{reduction in prediction error using model}}{\text{baseline prediction error}}$$

More on Model fit

- ▶ Baseline prediction error without a regression is using the mean of Y to predict. This is called the **Total sum of squares**:

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- ▶ Combining SSR and TSS from earlier we get the **coefficient of determination**, or R^2 :

$$R^2 = \frac{\text{TSS} - \text{SSR}}{\text{TSS}} = \frac{\text{how much smaller LS prediction errors are vs mean prediction error using the mean}}{\text{prediction error using the mean}}$$

- ▶ **Can be very misleading!** Does not guarantee linear model is ideal for given data.

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- ▶ **Out-of-sample fit:** how well your model predicts new data.
- ▶ **Overfitting:** OLS optimizes in-sample fit; may do poorly out of sample.

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 - ▶ Consider above expression
 - ▶ β_1 is the change in Y on average for a one unit increase in X , holding constant \mathbf{Z} (i.e., ceteris paribus)

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 - ▶ β_1 is the change in Y on average for a one unit increase in X , holding constant \mathbf{Z} (i.e., ceteris paribus)
 - ▶ **Statistical control** in a cross-sectional study.
- ▶ How do we estimate the coefficients? \rightsquigarrow The same way as before!
Minimize SSR.

$$\text{SSR} = \sum_{i=1}^n (\epsilon_i)^2 = \sum_{i=1}^n (Y_i - (\hat{\alpha} + \hat{\beta} X_i + \hat{\gamma} \mathbf{Z}_i))^2$$

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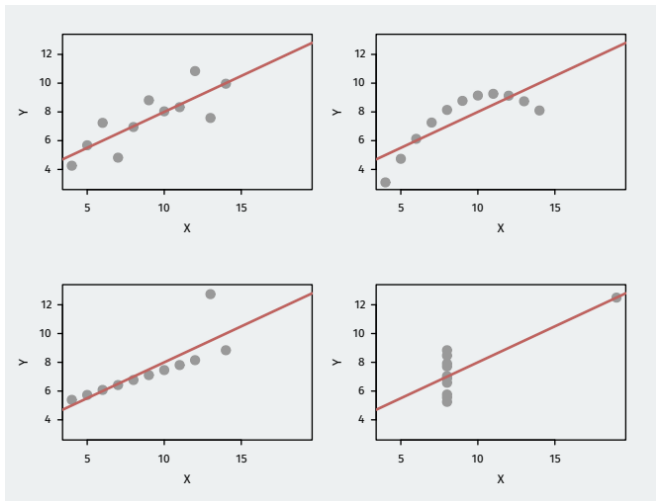
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 - ▶ But this could be overfitting!!
- ▶ Solution: penalize regression models with more variables.
 - ▶ Occam's razor: simpler models are preferred
- ▶ Adjusted R^2 : lowers regular R^2 for each additional covariate.
 - ▶ If the added covariates doesn't help predict, adjusted R^2 goes down!

Illustrating R-squared's Deficiencies

- ▶ What do you notice?

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- ▶ What do you notice? all these graphs have the same R^2



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$$\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 P_i + \hat{\beta}_2 T_i + \hat{\beta}_3 P_i \times T_i \text{ (Estimated)}$$

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- ▶ Note P_i is an indicator of period, and T_i is an indicator of whether the unit belongs to a group that is ever treated.
- ▶ Recall the DiD estimator takes the difference in the outcome among the treated group ($T_i = 1$) between $P_i = 1$ and $P_i = 0$ and subtracts the difference in control group ($T_i = 0$) between $P_i = 1$ and $P_i = 0$.

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DiD Estimator Proof

$$\frac{\text{Treated group } (T_i = 1) \text{ at } P_i = 1}{\text{Treated group } (T_i = 1) \text{ at } P_i = 0} \mid Y_i = \beta_0 + \beta_1 + \beta_2 + \beta_3 + \epsilon_i$$

DiD Estimator Proof

$$\begin{array}{l|l} \text{Treated group } (T_i = 1) \text{ at } P_i = 1 & Y_i = \beta_0 + \beta_1 + \beta_2 + \beta_3 + \epsilon_i \\ \hline \text{Treated group } (T_i = 1) \text{ at } P_i = 0 & Y_i = \beta_0 + \beta_2 + \epsilon_i \end{array}$$

DiD Estimator Proof

Treated group ($T_i = 1$) at $P_i = 1$	$Y_i = \beta_0 + \beta_1 + \beta_2 + \beta_3 + \epsilon_i$
Treated group ($T_i = 1$) at $P_i = 0$	$Y_i = \beta_0 + \beta_2 + \epsilon_i$
Control group ($T_i = 0$) at $P_i = 1$	

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Control group ($T_i = 0$) at $P_i = 1$	$Y_i = \beta_0 + \beta_1 + \epsilon_i$

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Control group ($T_i = 0$) at $P_i = 1$	$Y_i = \beta_0 + \beta_1 + \epsilon_i$
Control group ($T_i = 0$) at $P_i = 0$	

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Control group ($T_i = 0$) at $P_i = 0$	$Y_i = \beta_0 + \epsilon_i$

DiD Estimator Proof

Treated group ($T_i = 1$) at $P_i = 1$	$Y_i = \beta_0 + \beta_1 + \beta_2 + \beta_3 + \epsilon_i$
Treated group ($T_i = 1$) at $P_i = 0$	$Y_i = \beta_0 + \beta_2 + \epsilon_i$
Control group ($T_i = 0$) at $P_i = 1$	$Y_i = \beta_0 + \beta_1 + \epsilon_i$
Control group ($T_i = 0$) at $P_i = 0$	$Y_i = \beta_0 + \epsilon_i$

$$\begin{aligned} & (\text{Change in Treated Group}) - (\text{Change in Control Group}) = \\ & (\beta_0 + \beta_1 + \beta_2 + \beta_3 + \epsilon_i - (\beta_0 + \beta_2 + \epsilon_i)) - (\beta_0 + \beta_1 + \epsilon_i - (\beta_0 + \epsilon_i)) = \beta_3 \end{aligned}$$

DiD Estimator Proof

Treated group ($T_i = 1$) at $P_i = 1$	$Y_i = \beta_0 + \beta_1 + \beta_2 + \beta_3 + \epsilon_i$
Treated group ($T_i = 1$) at $P_i = 0$	$Y_i = \beta_0 + \beta_2 + \epsilon_i$
Control group ($T_i = 0$) at $P_i = 1$	$Y_i = \beta_0 + \beta_1 + \epsilon_i$
Control group ($T_i = 0$) at $P_i = 0$	$Y_i = \beta_0 + \epsilon_i$

$$\begin{aligned} & (\text{Change in Treated Group}) - (\text{Change in Control Group}) = \\ & (\beta_0 + \beta_1 + \beta_2 + \beta_3 + \epsilon_i - (\beta_0 + \beta_2 + \epsilon_i)) - (\beta_0 + \beta_1 + \epsilon_i - (\beta_0 + \epsilon_i)) = \beta_3 \end{aligned}$$

- Upshot: β_3 (coefficient on interaction of treatment group indicator and period indicator) is the DiD ATT estimator!

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```
lm(y ~ x + as.factor(unit_indicator))
```

OR

```
lm(y ~ x + as.factor(time_indicator))
```